

**A LOCAL VERSION OF THE KEMENEY CONSTANT
PHENOMENON AND RANDOM SEARCH OPTIMIZATION
FOR ONE-DIMENSIONAL DIFFUSIONS**

ROSS G. PINSKY

ABSTRACT. Consider a positive recurrent, one-dimensional diffusion process. Let τ_y denote the first time it hits y and denote by μ_{inv} and $\mu_{\text{inv}}(x)$ its invariant probability distribution and density. Consider the quantity $K_{\mu_{\text{inv}}}(x) := \int_{-\infty}^{\infty} (E_x \tau_y) \mu_{\text{inv}}(y) dy$, which is the expected time starting from $x \in \mathbb{R}$ until the diffusion process hits a point that is randomly chosen according to the invariant measure of the process. In [7], it was shown that this quantity is independent of x . This is a generalization of the classical Kemeny constant result for positive recurrent, irreducible finite state Markov chains. The fact that the above quantity is independent of x is only nontrivial if the quantity is finite, and it was also shown in [7] that the quantity is finite if and only if $\pm\infty$ are entrance boundaries for the diffusion. The entrance boundary requirement is a very strong restriction on a positive recurrent diffusion. In this paper, we prove a result of the same flavor that is nontrivial for all positive recurrent diffusions. Namely, let $[c, d]$ be an interval and let μ be any probability measure on \mathbb{R} that satisfies $\mu((-\infty, x]) = \mu_{\text{inv}}((-\infty, x])$, for $x \in [c, d]$. Let $K_{\mu}(x) = \int_{-\infty}^{\infty} (E_x \tau_y) \mu(dy)$. Then $K_{\mu}(x)$ is constant for $x \in [c, d]$. The measure μ can always be chosen so that the constant is finite; in particular, it will always be finite if μ is chosen to have compact support. We also study the local minima and maxima of K_{μ} for general probability measures μ without any reference to μ_{inv} . The quantity $K_{\mu}(x)$ may be thought of as the expected time for a diffusive searcher starting from x to locate a target that has been randomly positioned according to the distribution μ .

2010 *Mathematics Subject Classification.* 60C05, 05A05.

Key words and phrases. Kemeny constant, diffusion process, invariant measure, random search, positive recurrence .

1. INTRODUCTION AND STATEMENT OF RESULTS

Before defining a local version of the Kemeny constant for one-dimensional diffusion processes, we recall some results concerning the classical Kemeny constant. Let $\{X_n\}_{n=0}^\infty$ be an irreducible, discrete time Markov chain on a finite state space S , and denote its invariant probability measure by μ . For $j \in S$, let $\tau_j = \inf\{n \geq 0 : X_n = j\}$ denote the hitting time of j . Denoting expectations for the process starting from $i \in S$ by E_i , consider the quantity $\sum_{j \in S} \mu_j E_i \tau_j$, which is the expected time starting from the state i until the Markov chain hits a point that is randomly chosen according to the invariant measure of the process. In their book on Markov chains [4], Kemeny and Snell showed that the above quantity is independent of the initial state i ; this quantity has become known as *Kemeny's constant*. See also, the nice argument by Doyle [2].

In [1], the Kemeny constant phenomenon was studied for irreducible, positive recurrent, discrete time and continuous time Markov chains on a denumerably infinite state space S . It was shown that the quantity $\sum_{j \in S} \mu_j E_i \tau_j$ is either infinite for all $i \in S$, or else is finite and independent of i . The authors conjectured that this quantity is always infinite in the discrete time setting, and they proved this in the case of discrete time birth and death chains on $\{0, 1, \dots\}$. In the case of continuous time birth and death chains on $\{0, 1, \dots\}$, they proved that the Kemeny constant is finite if and only if ∞ is an *entrance boundary* for the process.

In [7], we considered the corresponding problem in the context of positive recurrent, one-dimensional Markov diffusion processes $X(t)$ on \mathbb{R} . The generator of $X(t)$ was assumed to be of the form

$$(1.1) \quad L = \frac{1}{2}a(x)\frac{d^2}{dx^2} + b(x)\frac{d}{dx},$$

where a is continuous and positive and b is locally bounded and measurable. Denote probabilities and expectations for the Markov process $X(\cdot)$ starting from $x \in \mathbb{R}$ by P_x and E_x . For $y \in \mathbb{R}$, let $\tau_y = \inf\{t \geq 0 : X(t) = y\}$ denote the first hitting time of y . As is well known [5], positive recurrence is equivalent to the condition

$$(1.2) \quad \int_{-\infty}^{\infty} \frac{1}{a(x)} \exp\left(2 \int_0^x \frac{b(t)}{a(t)} dt\right) dx < \infty,$$

and when the diffusion is positive recurrent, its invariant distribution μ_{inv} has density, which we call $\mu_{\text{inv}}(x)$, given by

$$(1.3) \quad \mu_{\text{inv}}(x) = c_0 \frac{1}{a(x)} \exp\left(2 \int_0^x \frac{b(t)}{a(t)} dt\right), \quad x \in \mathbb{R},$$

where c_0 is the normalization constant. The quantity

$$K_{\mu_{\text{inv}}}(x) := \int_{-\infty}^{\infty} (E_x \tau_y) \mu_{\text{inv}}(y) dy, \quad x \in \mathbb{R},$$

is the expected time starting from $x \in \mathbb{R}$ until the diffusion process hits a point that is randomly chosen according to the invariant measure of the process. In [7] it was proven that $K_{\mu_{\text{inv}}}(x)$ is independent of x .

However, this result is only interesting if this Kemeny constant is finite. It was shown in [7] that this Kemeny constant is finite if and only if $\pm\infty$ are entrance boundaries for the diffusion, analogous to the situation for continuous time birth and death processes mentioned above. For use in section 3, we note that the proof in [7] shows that $\int^{\infty} (E_x \tau_y) \mu_{\text{inv}}(y) dy$ is finite if and only if $+\infty$ is an entrance boundary, and $\int_{-\infty} (E_x \tau_y) \mu_{\text{inv}}(y) dy$ is finite if and only if $-\infty$ is an entrance boundary.

The boundary $+\infty$ is called an *entrance boundary* if $\lim_{x \rightarrow \infty} P_x(\tau_y < t) > 0$, for some $y \in \mathbb{R}$ and some $t > 0$ (or equivalently, for all $y \in \mathbb{R}$ and all $t > 0$). A parallel definition is given for the boundary $-\infty$. For the precise criterion for an entrance boundary, in terms of the coefficients a and b of the generator L of the process, see [7]. Heuristically, in order for $\pm\infty$ to be entrance boundaries, the drift $b(x)$ has to push the process inward sufficiently fast from $\pm\infty$. In particular, we note the following comparison between the requirement for positive recurrence and the requirement for an entrance boundary:

- (1.4)
- if $a(x)$ is bounded and bounded away from 0 and $b(x) = -\text{sgn}(x)c|x|^l$,
for large $|x|$, with $c > 0$ and $l \in \mathbb{R}$, then
 - i.* $\pm\infty$ are entrance boundaries if and only if $l > 1$;
 - ii.* the diffusion is positive recurrent if and only if either $l > -1$
or $l = -1$ and $c > \frac{1}{2}$.

We point out that if one considers a diffusion process on a finite interval, with reflection at the boundaries, in which case the diffusion is always positive recurrent, one can use the techniques in [7] to show that the corresponding Kemeny quantity is constant and always finite, analogous to the case of a Markov chain on a finite state space.

With the above background at hand, we turn to the problem studied in this paper. Let $y \in \mathbb{R}$ denote an unknown random target with known distribution $\mu \in \mathcal{P}(\mathbb{R})$, the space of probability measures on \mathbb{R} . A diffusive searcher $X(t)$, generated by L as above, starts out from some point $x \in \mathbb{R}$. The expected time to locate the target is given by

$$(1.5) \quad K_\mu(x) = \int_{-\infty}^{\infty} (E_x \tau_y) \mu(dy), \quad x \in \mathbb{R}.$$

We only consider distributions μ for which $K_\mu(x)$ is finite. (If it is finite for one x , then it is finite for all x .) It is then natural to consider where K_μ attains its minimum, and perhaps also, its maximum. Note, however, that it follows from the definition of an entrance boundary that if $+\infty$ ($-\infty$) is not an entrance boundary, then $\lim_{x \rightarrow +\infty} K_\mu(x) = \infty$ ($\lim_{x \rightarrow -\infty} K_\mu(x) = \infty$). Of course, if $\mu = \mu_{\text{inv}}$, which under our assumption that K_μ is finite is allowed only if $\pm\infty$ are entrance boundaries, then by the result from [7] noted above, $K_\mu(x)$ is constant.

Before continuing in the exposition, we note some background with regard to this random search problem. In [6], this setup was studied from a different angle; the diffusion coefficient a was assumed to be constant, the starting point of the searcher was fixed at $x = 0$, and the problem was to find for each target distribution μ , the drift b that minimizes $K_\mu(0)$. Note that if one lets the diffusive searcher be standard Brownian motion, then the process is not positive recurrent and consequently K_μ is infinite for all choices of non-degenerate probability measures μ . In order to overcome this issue, for a Brownian searcher one considers the problem with resetting; namely, at an exponential time, the searcher jumps back to its starting point and begins its search anew; this ensures the finiteness of the expected search time. This has been a much studied topic over the past 15 years; see [3] for a review article and many references.

Our interest in the minimum and maximum of K_μ led us to what we call a local version of the Kemeny constant phenomenon. Here is the result.

Theorem 1. *Let L , as in (1.1), be the generator of a positive recurrent diffusion on \mathbb{R} ; that is, assume (1.2) holds. Denote by μ_{inv} the invariant distribution with invariant density $\mu_{inv}(x)$ as in (1.3). Let $-\infty < c < d < \infty$. Let $\mu \in \mathcal{P}(\mathbb{R})$ be such that K_μ is finite, where K_μ is as in (1.5). Assume that*

$$(1.6) \quad \mu((-\infty, x]) = \mu_{inv}((-\infty, x]), \text{ for all } x \in [c, d].$$

Then $K_\mu(x)$ is constant for $x \in [c, d]$.

Remark 1. As noted above, the fact that $K_{\mu_{inv}}(x)$ is constant is only non-trivial if it is finite, and it is finite if and only if $\pm\infty$ are entrance boundaries for the diffusion process $X(t)$. This requirement is a very restrictive one on a positive recurrent diffusion, as (1.4) demonstrates. Theorem 1, on the other hand, is nontrivial for all positive recurrent diffusions. Indeed, fix c, d as in the theorem. Let ν_- and ν_+ be sub-probability measures, supported respectively on $(-\infty, c]$ and on $[d, \infty)$, that satisfy $\nu_-((-\infty, c]) = \mu_{inv}((-\infty, c])$ and $\nu_+([d, \infty)) = \mu_{inv}([d, \infty))$. Assume also that $\int_{-\infty}^{\infty} (E_x \tau_y) \nu_{\pm}(dy) < \infty$. (This last requirement is always fulfilled, for example, if ν_{\pm} are chosen to have compact support.) Then the theorem applies to the distribution μ defined by

$$\mu(A) = \begin{cases} \nu_-(A), & A \subset (-\infty, c]; \\ \mu_{inv}(A), & A \subset (c, d); \\ \nu_+(A), & A \subset [d, \infty). \end{cases}$$

Remark 2. If K_μ is finite and satisfies (1.6), then without reference to the conclusion of Theorem 1, we have

$$\begin{aligned}
(1.7) \quad K_\mu(x) &= \int_{-\infty}^c (E_x \tau_y) \mu(dy) + \int_c^d (E_x \tau_y) \mu_{\text{inv}}(y) dy + \int_d^\infty (E_x \tau_y) \mu(dy) = \\
&= \int_{-\infty}^c (E_x \tau_c + E_c \tau_y) \mu(dy) + \int_c^d (E_x \tau_y) \mu_{\text{inv}}(y) dy + \int_d^\infty (E_x \tau_d + E_d \tau_y) \mu(dy) = \\
&= \left[\mu_{\text{inv}}((-\infty, c]) E_x \tau_c + \int_c^d (E_x \tau_y) \mu_{\text{inv}}(y) dy + \mu_{\text{inv}}((d, \infty)) E_x \tau_d \right] + \\
&= \left[\int_{-\infty}^c (E_c \tau_y) \mu(dy) + \int_d^\infty (E_d \tau_y) \mu(dy) \right] = \\
&= F_1(x) + F_2(\mu), \text{ for } x \in [c, d],
\end{aligned}$$

where $F_1(x)$, which is independent of μ , denotes the third line in the above equation and $F_2(\mu)$, which is independent of x , denotes the fourth line in the above equation. This shows that for any two probability measures μ_1, μ_2 satisfying (1.6), and for which K_{μ_1} and K_{μ_2} are finite, the difference $K_{\mu_2}(x) - K_{\mu_1}(x)$ is constant over $x \in [c, d]$. Thus, (1.7) along with the Kemeny result in [7] prove Theorem 1 in the case that $K_{\mu_{\text{inv}}}$ is finite; indeed, let $\mu = \mu_1$ and $\mu_{\text{inv}} = \mu_2$. We also note that (1.7) along with Theorem 1 give the following corollary to Theorem 1.

Corollary 1. *Let $-\infty < c < d < \infty$. The quantity*

$$\mu_{\text{inv}}((-\infty, c]) E_x \tau_c + \int_c^d (E_x \tau_y) \mu_{\text{inv}}(y) dy + \mu_{\text{inv}}((d, \infty)) E_x \tau_d$$

is independent of $x \in [c, d]$.

We now return to the question of locating the local minima and maxima of $K_\mu(x)$ for general μ .

Theorem 2. *Let L as in (1.1) be the generator of a positive recurrent diffusion on \mathbb{R} ; that is, assume (1.2) holds. Denote by μ_{inv} the invariant distribution with invariant density $\mu_{\text{inv}}(x)$ as in (1.3). Let μ be a probability measure on \mathbb{R} for which $K_\mu(x)$, defined in (1.5), is finite.*

i.

$$(1.8) \quad K'_\mu(x) = \frac{2}{c_0} \exp\left(-2 \int_0^x \frac{b(t)}{a(t)} dt\right) (\mu((-\infty, x]) - \mu_{\text{inv}}((-\infty, x])),$$

where c_0 is as in (1.3). In particular, $K'_\mu(x) = 0$ if and only if $\mu((-\infty, x]) = \mu_{inv}((-\infty, x])$.

ii. Assume that μ possesses a continuous density $\mu(x)$. If $K'_\mu(x_0) = 0$, then x_0 is a local minimum for K_μ if $\mu(x_0) > \mu_{inv}(x_0)$, and x_0 is a local maximum for K_μ if $\mu(x_0) < \mu_{inv}(x_0)$. If $\mu(x_0) = \mu_{inv}(x_0)$, and $\mu(x) - \mu_{inv}(x)$ changes sign at $x = x_0$, then x_0 is neither a maximum nor a minimum for K_μ , while if $\mu(x) - \mu_{inv}(x)$ is positive (negative) for $|x - x_0|$ sufficiently small, then x_0 is a minimum (maximum).

Note that Theorem 1 follows from part (i) of Theorem 2. We prove Theorem 2 in section 2. In section 3 we illustrate Theorem 2 with some examples.

2. PROOF OF THEOREM 2

Proof of part (i). We have the following explicit expression for the expected hitting time:

$$(2.1) \quad E_x \tau_y = \begin{cases} 2 \int_y^x dz \exp(-2 \int_0^z \frac{b(t)}{a(t)} dt) \int_z^\infty dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt), & -\infty < y < x; \\ 2 \int_x^y dz \exp(-2 \int_0^z \frac{b(t)}{a(t)} dt) \int_{-\infty}^z dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt), & x < y < \infty. \end{cases}$$

For a derivation, see for example the proof of Proposition 2 in [6] (where $a(x)$ is a constant and denoted by D). Thus,

$$(2.2) \quad K_\mu(x) = \int_{-\infty}^x \left(2 \int_y^x dz \exp(-2 \int_0^z \frac{b(t)}{a(t)} dt) \int_z^\infty dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt) \right) \mu(dy) + \int_x^\infty \left(2 \int_x^y dz \exp(-2 \int_0^z \frac{b(t)}{a(t)} dt) \int_{-\infty}^z dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt) \right) \mu(dy).$$

Denote the first summand in (2.2) by $I(x)$ and the second summand there by $II(x)$. Denote the integrand in $I(x)$ by $G(x, y)$, so $I(x) = \int_{-\infty}^x G(x, y) \mu(dy)$. Recall that μ is an arbitrary probability measure on \mathbb{R} . However, since $G(x, x) = 0$, we do not need to be careful about the possibility of μ having an atom at the upper limit x of the integral. For $h > 0$, We write

$$(2.3) \quad \frac{1}{h} (I(x+h) - I(x)) = \int_{-\infty}^{x+h} \frac{1}{h} (G(x+h, y) - G(x, y)) \mu(dy) + \frac{1}{h} \int_x^{x+h} G(x, y) \mu(dy).$$

Using the mean value theorem, we write the first summand on the right hand side of (2.3) as

$$(2.4) \quad \int_{-\infty}^{x+h} \frac{1}{h} (G(x+h, y) - G(x, y)) \mu(dy) = \int_{-\infty}^{x+h} \frac{\partial G}{\partial x}(x + h^*(x, y), y) \mu(dy), \text{ where } 0 < h^*(x, y) < h.$$

Note that $\frac{\partial G}{\partial x}(z, y) = 2 \exp(-2 \int_0^z \frac{b(t)}{a(t)} dt) \int_z^\infty dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt)$, and is thus independent of y . Thus, from (2.4) we have

$$(2.5) \quad \lim_{h \rightarrow 0^+} \int_{-\infty}^{x+h} \frac{1}{h} (G(x+h, y) - G(x, y)) \mu(dy) = \left(2 \exp(-2 \int_0^x \frac{b(t)}{a(t)} dt) \int_x^\infty dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt) \right) \mu((-\infty, x]).$$

Since $G(x, y)$ is continuously differentiable in y , and since $G(x, x) = 0$, we have $|G(x, y)| = O(|y - x|)$. Using this estimate and again the fact that $G(x, x) = 0$, the second summand on the right hand side of (2.3) satisfies

$$(2.6) \quad \left| \frac{1}{h} \int_x^{x+h} G(x, y) \mu(dy) \right| = \frac{O(h)}{h} \mu((x, x+h]).$$

From (2.6) it follows that

$$(2.7) \quad \lim_{h \rightarrow 0^+} \frac{1}{h} \int_x^{x+h} G(x, y) \mu(dy) = 0.$$

We conclude from (2.3), (2.5) and (2.7) along with similar calculations when $h \rightarrow 0^-$ that

$$(2.8) \quad \lim_{h \rightarrow 0} \frac{1}{h} (I(x+h) - I(x)) = \left(2 \exp(-2 \int_0^x \frac{b(t)}{a(t)} dt) \int_x^\infty dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt) \right) \mu((-\infty, x]).$$

A similar calculation for $II(x)$ shows that

$$(2.9) \quad \lim_{h \rightarrow 0} \frac{1}{h} (II(x+h) - II(x)) = - \left(2 \exp(-2 \int_0^x \frac{b(t)}{a(t)} dt) \int_{-\infty}^x dw \frac{1}{a(w)} \exp(2 \int_0^w \frac{b(t)}{a(t)} dt) \right) \mu([x, \infty)).$$

From (2.3), (2.8) and (2.9), we obtain

$$(2.10) \quad K'_\mu(x) = \left(2 \exp\left(-2 \int_0^x \frac{b(t)}{a(t)} dt\right) \int_x^\infty dw \frac{1}{a(w)} \exp\left(2 \int_0^w \frac{b(t)}{a(t)} dt\right) \right) \mu((-\infty, x]) - \left(2 \exp\left(-2 \int_0^x \frac{b(t)}{a(t)} dt\right) \int_{-\infty}^x dw \frac{1}{a(w)} \exp\left(2 \int_0^w \frac{b(t)}{a(t)} dt\right) \right) \mu([x, \infty)).$$

Recalling the formula in (1.3) for the invariant density $\mu_{\text{inv}}(x)$, we conclude from (2.10) that

$$(2.11) \quad K'_\mu(x) = \frac{2}{c_0} \exp\left(-2 \int_0^x \frac{b(t)}{a(t)} dt\right) (\mu_{\text{inv}}([x, \infty))\mu((-\infty, x]) - \mu_{\text{inv}}((-\infty, x])\mu([x, \infty))).$$

The right hand side of (2.11) is equivalent to the right hand side of (1.8).

Proof of part (ii). We are now assuming that μ has a continuous density $\mu(x)$. Differentiating (1.8), we have

$$(2.12) \quad K''_\mu(x) = \frac{2}{c_0} \exp\left(-2 \int_0^x \frac{b(t)}{a(t)} dt\right) \times \left(-2 \frac{b(x)}{a(x)} (\mu((-\infty, x]) - \mu_{\text{inv}}((-\infty, x])) + \mu(x) - \mu_{\text{inv}}(x) \right).$$

From (1.8), it follows that if $K'_\mu(x_0) = 0$, then the sign of $K''_\mu(x_0)$ is the same as that of $\mu(x_0) - \mu_{\text{inv}}(x_0)$. This proves part (ii) except for the case that $\mu(x_0) = \mu_{\text{inv}}(x_0)$. Consider now the case that $\mu(x_0) = \mu_{\text{inv}}(x_0)$. If $\mu(x) > \mu_{\text{inv}}(x)$ when $x \neq x_0$ and $|x - x_0|$ is sufficiently small, then by (1.8), it follows that for $x \neq x_0$ and $|x - x_0|$ is sufficiently small, $K'_\mu(x)$ is negative when $x < x_0$ and positive when $x > x_0$, and consequently x_0 is a minimum for K_μ . If $\mu(x) < \mu_{\text{inv}}(x)$ when $x \neq x_0$ and $|x - x_0|$ is sufficiently small, then similar reasoning shows that x_0 is a maximum for K_μ . Finally, if $\mu(x) - \mu_{\text{inv}}(x)$ changes sign at x_0 , then similar reasoning shows that x_0 is neither a maximum nor a minimum for K_μ . \square

3. SOME EXAMPLES

We illustrate Theorem 2 by considering a generic set of examples and then concretize this with specific examples. Our generic assumption is that μ has a density $\mu(x)$ and the graphs of the densities $\mu_{\text{inv}}(x)$ and $\mu(x)$ intersect one another a finite number of times. It then follows from Theorem 2 that

K_μ is monotone in a neighborhood of $+\infty$ and in a neighborhood of $-\infty$; consequently, $K_\mu(\pm\infty) = \lim_{x \rightarrow \pm\infty} K_\mu(x) \in (0, \infty]$ exists.

We divide the above assumption into the following three situations:

- I.** $\mu(x) < \mu_{\text{inv}}(x)$ for sufficiently large $|x|$;
- II.** $\mu_{\text{inv}}(x) < \mu(x)$ for sufficiently large $|x|$;
- III-a.** $\mu(x) < \mu_{\text{inv}}(x)$ for x in a neighborhood of $+\infty$ and $\mu_{\text{inv}}(x) < \mu(x)$ for x in a neighborhood of $-\infty$;
- III-b.** $\mu_{\text{inv}}(x) < \mu(x)$ for x in a neighborhood of $+\infty$ and $\mu(x) < \mu_{\text{inv}}(x)$ for x in a neighborhood of $-\infty$.

As in the theorem, we always assume that $K_\mu(x) < \infty$, for $x \in \mathbb{R}$. Thus by the theorem from [7] and its proof, as noted in section 1, if neither of $\pm\infty$ is an entrance boundary, then μ must satisfy (I). If $+\infty$ is an entrance boundary and $-\infty$ is not one, then μ must satisfy either (I) or (III-b), while if $+\infty$ is not an entrance boundary and $-\infty$ is one, then μ must satisfy either (I) or (III-a). If both $\pm\infty$ are entrance boundaries, then for any of the three conditions above, there exist appropriate μ satisfying that condition. By straight forward analysis, which we omit, Theorem 2 allows for the following conclusions under each of the above situations:

Theorem 2 under condition (I). Then $|\{x \in \mathbb{R} : \mu(x) = \mu_{\text{inv}}(x)\}| = 2n$, for some $n \in \mathbb{N}$. The function $K_\mu(x)$, $x \in \mathbb{R}$, possesses k local minima and $k - 1$ local maxima, for some $k \in [n]$. In addition, $\pm\infty$ are local maxima. In particular, if $+\infty$ ($-\infty$) is not an entrance boundary, then $K_\mu(+\infty) = \infty$ ($K_\mu(-\infty) = \infty$). If $k = 1$, then then the unique local minimum is the global minimum for K_μ , and the supremum is at $+\infty$, $-\infty$ or both.

Theorem 2 under condition (II). Since we are assuming that $K_\mu < \infty$, in this case $\pm\infty$ are necessarily entrance boundaries. We have $|\{x \in \mathbb{R} : \mu(x) = \mu_{\text{inv}}(x)\}| = 2n$, for some $n \in \mathbb{N}$. The function $K_\mu(x)$, $x \in \mathbb{R}$, possesses $k - 1$ local minima and k local maxima, for some $k \in [n]$. Also $\pm\infty$ are local minima. If $k = 1$, then then the unique local maximum is the global maximum for K_μ , and the infimum is at $+\infty$, $-\infty$ or both.

Theorem 2 under condition (III-a). Since we are assuming that $K_\mu < \infty$, in this case necessarily $-\infty$ is an entrance boundary. We have $|\{x \in \mathbb{R} : \mu(x) = \mu_{\text{inv}}(x)\}| = 2n - 1$, for some $n \in \mathbb{N}$. The function $K_\mu(x)$ possesses k local minima and k local maxima, for some $k \in \{0, \dots, n - 1\}$. Also, $-\infty$ is

a local minimum and $+\infty$ is a local maximum. In particular, if $+\infty$ is not an entrance boundary, then $K_\mu(+\infty) = \infty$. If $k = 0$, then the infimum of K_μ is at $-\infty$ and the supremum of K_μ is at $+\infty$.

Theorem 2 under condition (III-b). The description above for condition (III-a) holds with the roles of $+\infty$ and $-\infty$ interchanged.

We now consider three concrete examples. For the diffusion process, we assume that the diffusion coefficient is $a(x) = 1$ and the drift is $b(x) = -\text{sgn}(x)\beta|x|^l$, with $l > -1$ and $\beta > 0$. By (1.4), the diffusion is positive recurrent; the boundaries $\pm\infty$ are entrance boundaries if and only if $l > 1$. Define

$$(3.1) \quad N(c, m) = \left(\int_{-\infty}^{\infty} e^{-c|w|^m} dw \right)^{-1}, \quad c, m > 0.$$

By (1.3), the invariant density is given by

$$(3.2) \quad \mu_{\text{inv}}(x) = N\left(\frac{2\beta}{l+1}, l+1\right) e^{-\frac{2\beta}{l+1}|x|^{l+1}}, \quad x \in \mathbb{R}.$$

We have the following proposition.

Proposition 1. *Let $a(x) = 1$ and $b(x) = -\text{sgn}(x)\beta|x|^l$, with $l > -1$ and $\beta > 0$. Then for all $x \in \mathbb{R}$,*

$$(3.3) \quad E_x \tau_y \sim \frac{1}{\beta N\left(\frac{2\beta}{l+1}, l+1\right)} |y|^{-l} e^{\frac{2\beta}{l+1}|y|^{l+1}}, \quad \text{as } |y| \rightarrow \infty.$$

Remark. Note from (3.2) and (3.3) that K_μ is finite if and only if $l > 1$, and by (1.4), $\pm\infty$ are entrance boundaries if and only if $l > 1$, verifying in this particular case what the general result from [7] guarantees.

Proof of Proposition 1. By symmetry, it suffices to consider the case $y \rightarrow +\infty$. From (2.1), it is clear that with $a(x)$ and $b(x)$ as in the statement of the proposition, $\lim_{y \rightarrow +\infty} E_x \tau_y = \infty$, and consequently the leading order asymptotic behavior of $E_x \tau_y$ as $y \rightarrow +\infty$ is independent of x . Thus, we can assume that $x > 0$, in which case we have for $x < y$,

$$(3.4) \quad E_x \tau_y = 2 \int_x^y dz \exp\left(\frac{2\beta z^{l+1}}{l+1}\right) \int_{-\infty}^z dw \exp\left(-\frac{2\beta|w|^{l+1}}{l+1}\right) \sim \frac{2}{N\left(\frac{2\beta}{l+1}, l+1\right)} \int_x^y \exp\left(\frac{2\beta z^{l+1}}{l+1}\right) dz, \quad \text{as } y \rightarrow +\infty.$$

Using l'Hôpital's rule at ∞ shows that

$$(3.5) \quad \lim_{y \rightarrow +\infty} \frac{\int_x^y \exp(\frac{2\beta z^{l+1}}{l+1}) dz}{(2\beta y^l)^{-1} \exp(\frac{2\beta y^{l+1}}{l+1})} = 1.$$

From (3.4) and (3.5), we obtain (3.3) for $y \rightarrow +\infty$. \square

For convenience, in what follows we will refer to the above diffusion process as a (β, l) -diffusion. Here now are three concrete examples.

1. (β, l) -diffusion, with $l > -1$, and with target density

$$\mu(x) = N(c, l+1)e^{-c|x|^{l+1}}, \text{ with } c > \frac{2\beta}{l+1}:$$

It follows from Proposition 1 that K_μ is finite. The diffusion and the target distribution correspond to condition (I) above. Clearly, $|\{x \in \mathbb{R} : \mu(x) = \mu_{\text{inv}}(x)\}| = 2$. So by Theorem 2 applied to condition (I), we conclude that K_μ possesses a unique local minimum and that $\pm\infty$ are local maxima. By the symmetry, $\mu((-\infty, 0]) = \mu_{\text{inv}}((-\infty, 0]) = \frac{1}{2}$. Thus the unique local minimum of K_μ is at $x = 0$, and this is the global minimum. By symmetry, the supremum occurs at both $\pm\infty$. In particular, if $l \in (-1, 1]$, then $\pm\infty$ are not entrance boundaries, so $K_\mu(\pm\infty) = \infty$.

2. (β, l) diffusion, with $l > -1$, and with target density $\mu(x) = N(c, m)e^{-c|x|^m}$, with $m > l+1$ and $c > 0$:

It follows from Proposition 1 that K_μ is finite. The diffusion and the target distribution correspond to condition (I) above. Thus, by Theorem 2 applied to condition (I), $\pm\infty$ are local maxima for K_μ . In particular, if $l \in (-1, 1]$, then $\pm\infty$ are not entrance boundaries, so $K_\mu(\pm\infty) = \infty$. Clearly, there exists a $c_0 = c_0(\beta) > 0$ such that

$$(3.6) \quad \begin{aligned} \mu(0) = N(c, m) < N(\frac{2\beta}{l+1}, l+1) = \mu_{\text{inv}}(0), & \text{ if } c < c_0; \\ \mu(0) = N(c, m) > N(\frac{2\beta}{l+1}, l+1) = \mu_{\text{inv}}(0), & \text{ if } c > c_0; \end{aligned}$$

Thus, by symmetry and Theorem 2, K_μ has a maximum at $x = 0$ if $c < c_0$ and a minimum at $x = 0$ if $c > c_0$. For $c = c_0$, we have $\mu(0) = \mu_{\text{inv}}(0)$ and $\mu(x) - \mu_{\text{inv}}(x)$ is positive for $|x|$ sufficiently small, so by Theorem 2, K_μ has a minimum at $x = 0$ if $c = c_0$. The equation $\mu(x) = \mu_{\text{inv}}(x)$, for $x > 0$ is equivalent to $cx^m - \frac{2\beta}{l+1}x^{l+1} = \log \frac{N(c, m)}{N(\frac{2\beta}{l+1}, l+1)}$, for $x > 0$. The function $cx^m - \frac{2\beta}{l+1}x^{l+1}$ on $[0, \infty)$ vanishes at $x = 0$ and it has a unique minimum, which is negative; it decreases to the left of that minimum and

increases unboundedly to the right of that minimum. Note from (3.6) that $\log \frac{N(c,m)}{N(\frac{2\beta}{l+1}, l+1)}$ is nonnegative if $c \geq c_0$ and negative if $c < c_0$. Thus, if $c \geq c_0$ it follows from the equation above that there is exactly one positive x for which $\mu(x) = \mu_{\text{inv}}(x)$. Consequently by Theorem 2 applied to condition (I) and by symmetry, K_μ has its unique local minimum at $x = 0$, which is the global minimum, and its supremum occurs at $\pm\infty$. On the other hand, for $c < c_0$, it follows from the equation above that there are either zero or two positive values of x for which $\mu(x) = \mu_{\text{inv}}(x)$. However, since both 0 and $\pm\infty$ are local maxima, it follows that K_μ must have a local minimum on $(0, \infty)$. Consequently, by Theorem 2 there must be two positive values of x for which $\mu(x) = \mu_{\text{inv}}(x)$. We conclude from Theorem 2 that K_μ has local maxima at 0 and $\pm\infty$, and two local minimum, which are the global minimum, at $\pm x_0$, where x_0 is the unique positive number which satisfies $\mu((-\infty, x_0]) = \mu_{\text{inv}}((-\infty, x_0])$.

3. (β, l) diffusion, with $l > 1$, and with target density

$\mu(x) = N_{a,\beta,l} \min(1, |x|^a) e^{-\frac{2\beta}{l+1}|x|^{l+1}}$, where $a \in (0, l-1)$ and $N_{a,\beta,l}$ is the appropriate normalization constant:

It follows from Proposition 1 that K_μ is finite. The diffusion and the target distribution correspond to condition (II) above. By Theorem 2 applied to condition (II), $\pm\infty$ are local minima. Clearly, there is a unique positive x for which $\mu(x) = \mu_{\text{inv}}(x)$. Consequently, by symmetry and Theorem 2 applied to condition (II), we conclude that K_μ attains a unique local maximum, which is the global maximum, at $x = 0$ and its infimum occurs at $\pm\infty$.

REFERENCES

- [1] Bini, D., Hunter, J., Latouche, G., Meini, B. and Taylor, P., *Why is Kemeny's Constant a Constant?*, J. Appl. Prob. **55** (2018), 1025-1036.
- [2] Doyle, P. G. (2009), *The Kemeny constant of a Markov chain*, preprint. Available at <http://arxiv.org/abs/0909.2636>.
- [3] Evans, M.R., Majumdar, S.N. and Schehr, G., *Stochastic resetting and applications*, J. Phys. A **53** (2020), no. 19, 193001, 67.
- [4] Kemeny, J. and Snell, J. L., *Finite Markov chains*, Reprinting of the 1960 original, Springer-Verlag, New York-Heidelberg, 1976.
- [5] Pinsky, R. G., *Positive Harmonic Functions and Diffusion*, Cambridge Studies in Advanced Mathematics **45**, Cambridge University Press, (1995).

- [6] Pinsky, R.G., *Optimizing the drift in a diffusive search for a random stationary target*, Electron. J. Probab. 24 (2019), Paper No. 82, 22 pp.
- [7] Pinsky, R.G., *Kemeny's constant for one-dimensional diffusions*, Electron. Commun. Probab. **24** (2019), Paper No. 36, 5 pp.

DEPARTMENT OF MATHEMATICS, TECHNION—ISRAEL INSTITUTE OF TECHNOLOGY,
HAIFA, 32000, ISRAEL

Email address: `pinsky@technion.ac.il`

URL: <https://pinsky.net.technion.ac.il/>